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OVER-QUALIFICATION AND DIMENSIONS OF JOB SATISFACTION

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Abstract

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JEL classification: D82, I26, J24, J28, J62

Keywords: over-qualification, domains of job satisfaction, Spain

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Abstract

Over-qualification may arise from voluntary decisions of individuals to acquire more qualifications than those required at the job place. In these cases, mismatch may have a role allowing workers to compensate the lack of some other skills, or to gain access to the labour market. Consequently, workers may feel no less satisfied, at least in some domains, than adequately matched workers. The aim of this paper is to analyse the relationship between over-qualification and different domains of job satisfaction in Spain, a country characterised by a strongly segmented labour market with high unemployment level, and a large number of mismatched. Using micro data for a representative sample of Spanish workers, we employ an IV estimation procedure to control for potential endogeneity arising from reverse causation or unobserved heterogeneity. Results obtained provide apparent evidence on that mismatched workers do not necessarily feel less satisfied than adequately matched workers in all domains of job satisfaction.

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1. Introduction

The generalisation of the phenomenon of over-qualification, which encompasses over-education, over-experience and skill under-utilisation, follows that of under-employment, a broader concept within which over-qualification is embraced (Feldman, 1996; Johnson et al., 2002). The topic of over-qualification has been extensively studied in the economic literature, with emphasis on the effects on returns and on job satisfaction (see the meta-analyses by Groot and Maassen van den Brink, 2000; Rubb, 2003) and on the phenomena of the bumping down and/or crowding out of less skilled workers (Hartog 2000; Lene, 2011; McGuinness, 2006).

When focusing on the association between over-qualification and job satisfaction a typical finding is a negative relationship, which can be explained under a psychological view by relative deprivation. According to this, when an individual wants an object and feel he/she deserves to get it but do not, that is, when a gap between aspirations and actual realisations exist, he/she becomes frustrated (Crosby 1976, 1984). Wages, responsibilities at job, challenges and career advancements are generally lower for the over-qualified (Bills, 1992; Khan and Morrow, 1991; Maynard and Hakel, 1999; Feldman et al., 2002; Johnson and Johnson, 2000; Johnson et al., 2002) causing a sense of deprivation and unfairness which is reflected in individuals showing lower job performance and more willingness to leave than the adequately matched (Bolino and Feldman, 2000; Feldman et al., 2002; Lee et al., 2016; Tsang et al., 1991; Verhaest and Omey, 2006, 2009). Under this view, firms would prefer hiring adequately qualified job applicants to those who are over-qualified (Bills, 1992; Maynard and Hakel, 1999).

Several findings challenge the relative deprivation theory. First, even though workers may feel unsatisfied because of over-qualification they do not necessarily perform worse than less-qualified, perfectly-matched workers (Holton et al., 2002; Fine and Nevo, 2008; Erdogan and Bauer, 2009). Thus, workers can stay in jobs for which they are over-qualified because these jobs have other good characteristics (McGuinness and Sloane, 2011). Additionally, firms may find over-education useful in assessing the ranking of a particular individual on the ability spectrum (Green et al. 2002), or as an indication of adaptive capacities (Lene 2011).

Second, using wage regressions, it is habitually found that the over-qualified obtain lower wages than equally qualified workers who are in properly matched jobs, but higher wages than less qualified workers in a well-matched job (Groot and Massen van den Brink, 2000; Sloane, 2003; McGuinness, 2006; Leuven and Oosterbeek, 2011; García and Montuenga, 2017). This result suggests that over-qualified are penalised

against the equally-qualified, properly-matched but extra-rewarded with respect to the well-matched, but less-qualified workers.

In consequence, it is not unsurprising to find studies disclosing the absence of relationship between over-education and job satisfaction (e.g. Groot and Maassen van den Brink, 2000; Büchel, 2002) and even a positive relationship between over-education and firm's productivity (Kampelman and Ryck, 2012). A first explanation for these results is that some determinants may moderate the negative relationship between job satisfaction and over-qualification. Thus, gender, marital status, emotional support or job-related conditions such as empowerment, salary or work experience may alleviate the negative influence of over-qualification on satisfaction (Erdogan and Bauer, 2009; Erdogan et al., 2011; Peiro et al., 2010; Lee et al., 2016). Additionally, given that job satisfaction can be seen as an array of different domains/facets of satisfaction related to performance in the job place, it may be the case that over-qualified workers may feel less satisfied in some domains and more satisfied in others than adequately matched, so that the final component may be of any sign (Johnson and Johnson, 2000; Taber and Alliger, 1995). For example, if workers over-qualify to obtain a permanent position or just to escape from unemployment, then they may report a valuation of satisfaction with job stability or job promotion not very different from that reported by an adequately matched worker, counterbalancing lower valuations corresponding to other domains, such as pay or self-fulfilment.

In this paper, we deepen into this argument by hypothesising that, individuals who over-qualify may be not less satisfied than adequately matched if, for example, they aim to mask or compensate their lack of skills or to access to employment. Whereas over-qualification is habitually considered as suboptimal, consequence of a mismatch due to search or job frictions, the increasing dispersion in ability and/or skills among equally educated workers may induce individuals to voluntarily acquire more qualifications than those they can productively use in their jobs in order to compensate for a lack in those observed and unobserved skills or as a way of gaining access to employment or to search for a better job (McGuinness and Wooden, 2009; Verhaest and Omey, 2009; McGuinness and Sloane, 2011; Garcia-Mainar and Montuenga-Gomez, 2017). Even more, this signalling role of over-education may be especially important in periods of recession, and/or in areas where unemployment is high, since the decision to invest in education is not only regarded as a way to have access to higher wages but also as a way of insuring against unemployment. Thus, Clark and Oswald (2002) and Verhaest and

Omey (2009) show that being unemployed produces a greater stigma than working at a job for which the individual is over-qualified.

We investigate the relationship between educational/skill mismatch and up to sixteen domains of job satisfaction for the Spanish case, a country with some specific features. First, in international perspective, the proportion of over-educated workers is among the highest within the OECD countries (OECD, 2011; Verhaest and van der Velden, 2013). Second, labour market is very slack; unemployment rate has consistently been among the highest within the EU during decades, and has risen to values over 25% during the Great Recession. Nowadays, it has reduced to values below 20%, but still doubling the OECD average. Third, labour market is strongly segmented, with a temporary rate that has been the highest within the EU, around 30%, during the last 30 years, generating a dual market and favouring the volatility of employment over time (Bentolila et al., 2012).

The data used in the paper, the Spanish Quality of Work Life Survey (*Encuesta de Calidad de Vida en el Trabajo*, ECVT henceforth), is of cross-sectional nature and prevents from longitudinal analysis. Hence, the use of Instrumental Variable (IV) estimation is needed in order to deal with endogeneity bias. In particular, the analysis of the association between each facet of job satisfaction and over-qualification is carried out through regression analysis taking into account both, the ordered nature of the dependent variable (each satisfaction domain) and the (plausible) appearance of endogeneity bias. Since satisfaction and over-qualification variables are both self-perceived by individuals, unobserved elements or reverse causation may be driving the final valuations of individuals. Results show that when endogeneity is not considered, over-qualification is systematically found to be negatively correlated with any domain of job satisfaction, corroborating the findings of previous literature. However, when dealing with endogeneity, in many dimensions of job satisfaction, the negative relationship is no longer observed.

The structure of the paper is as follows. Section 2 presents a summary of the literature and Section 3 describes the data, the diverse domains of job satisfaction and the concepts of mismatch used. In Section 4, we describe the methodology applied and present the estimated results. Finally, Section 5 concludes.

2. Literature review

Several models have been proposed in the literature as possible explanations for the existence of over-education. The traditional economic argument based on the human capital theory is recurrently unsupported by empirical evidence, as is also the Thurow's (1975) job competition theory (see Dolton and Vignoles, 2000; Rubb, 2003; McGuinness, 2006; Brynin and Longhi, 2009; McGuinness and Wooden, 2009; Baert et al., 2013). Whereas the relationship between job satisfaction and over-qualification is habitually shown to be negative, empirical studies based on wage regressions typically find that over-educated workers earn more than adequately matched workers in the same kind of jobs, but less than adequately matched workers with the same amount of education. In other terms, over-educated workers face a wage penalty compared to equally educated individuals who are job-matched. These results give support to an assignment theory (Sattinger, 1993; Sloane, 2003; McGuinness, 2006), which rests on that not all similarly educated workers are equally productive in all jobs, but there exists heterogeneous skill/ability distribution, provoking mismatch in the allocation of workers to jobs (Chevalier, 2003; Green and McIntosh, 2007). Thus, some authors have found the absence of relationship between over-education and job satisfaction (e.g. Groot and Maassen van den Brink, 2000; Büchel, 2002) and even a positive relationship between over-education and firm's productivity (Kampelman and Ryck, 2012). Sloane (2003) argues that, unlike educational mismatches, skill mismatches have a strong negative impact on job satisfaction (see also, Allen and van der Velden, 2001; and Green and Zhu, 2010). This result is observed also for Spain in Badillo-Amador and Vila (2013). These results stress the need of studying the more general concept of over-qualification instead of over-education.

Mismatch may be of transitory character if it disappears in a course of upwards career mobility (Sicherman and Galor, 1990; Sicherman, 1991; Alba-Ramírez, 1993) and then, over-qualified may not feel continuously less job-satisfied than adequately-matched. However, most of recent studies tend to challenge the career mobility hypothesis (see Baert et al., 2013 and references therein) and find that over-qualification is a permanent phenomenon, making possible over-qualified being not as job-satisfied as adequately matched, establishing hence a negative relationship between over-qualification and job satisfaction. Psychological literature usually finds that relative deprivation may explain why over-qualified workers feel dissatisfied and then perform worse than adequately matched and are more prone to leave the job (Bolino and

Feldman, 2000; Feldman et al., 2002; Khan and Morrow, 1991; Maynard et al., 2006; Peiro et al., 2010).

Several studies find that apart from initial mismatch due to job search frictions, or inefficiency in the labour market, over-qualification may appear by other reasons such as to compensate for deficient human capital in other respects, for the lack of other skills, to reveal they are indeed qualified for a job, or to disguise among other more able individuals (Green et al., 2002; Ordine and Rose, 2009; Chevalier and Lindley, 2009). In slack labour markets, where unemployment is large, over-qualification may be used also by individuals to either gain access to the labour market, to improve their position in wage bargaining, or to show adaptability to a changing environment in the job market (Charlot et al., 2005; Fernández, 2006; Lene, 2011). Ortiz (2010) finds that over-education is more common among permanent workers in Spain, since over-education allows workers not to achieve a better match, but a secure job. All these arguments suggest that over-qualification may be result from voluntary choices of individuals, and then not totally sub-optimal, thereby allowing for a non-negative relationship between over-qualification and job satisfaction.

Different studies, following the classification of Khan and Morro (1991), have differentiated between perceived mismatch (when education, experience or ability is higher than required in the job) and perceived no-grow (when the job hold by a workers lacks challenge), finding that the relationship with job satisfaction varies across dimensions. Johnson and Johnson (2000) further consider four facets of job satisfaction (overall, satisfaction with promotion, with supervisors and with pay). With surveyed data from 116 Postal Workers of a midwestern American Union local, they find at the cross sectional level that perceived no-grow is negatively related with work satisfaction and with satisfaction with supervision; and that perceived mismatch is negatively related with satisfaction with promotion and with satisfaction with pay. While it is found a negative relationship between the general domain of job satisfaction and over-qualification, this relationship is influenced by different factors, such as challenge, status, autonomy, etc., among which some of them may have positive consequences for over-qualified workers. They conclude that “job setting is multidimensional and composed of different constituent parts with which an individual may be either satisfied or dissatisfied” (Johnson and Johnson, 2000: p.552). Johnson et al. (2002) obtain similar conclusions when adding some indicators on the willingness to remain or not in the organisation.

Fine and Nevo (2008), and Erdogan and Bauer (2009), although obtain a negative relationship between perceived over-qualification and job satisfaction, provide convincing evidence of over-qualification to be positively related to job performance. Over-qualified are less satisfied than equally-qualified, adequately-matched, but they perform better than less-qualified, adequately-matched, showing that there are advantages to hiring employees who perceive that they are over-qualified.

Others authors have focused only on perceived mismatch since perceived no grow actually seems to allude to task characteristics (e.g., job interest, challenges), rather than to an education-related construct, and thus less adequate to capture over-qualification (Maynard et al., 2006; Peiro et al., 2010). Peiro et al. (2010) study the relationship between over-qualification and job satisfaction in Spain constructing three facets of job satisfaction: extrinsic, intrinsic and social significance. Extrinsic facets include salary, physical conditions, generous holidays, job security, promotion and working times. Intrinsic facets covers the facets about autonomy, skill utilization, task variety, learning opportunity, task significance, allow initiative and work with responsibilities. Finally, social significance includes contact with customer, contact with co-workers, social service, social status and supervisor guidance. All the individual facets are loaded into the three dimensions described through confirmatory factor analysis. They use a representative sample of 643 young Spanish employees (between 16 and 30 years old) living in the Valencian community and in the metropolitan cities of Barcelona and Madrid (Spain) to study the relationships between over-education and the three facets of job satisfaction as well as the direct and moderating role of salary, contract of employment, and work experience. After controlling for variables such as gender, age, education and region, they find a negative relationship between over-qualification and each of the three facets of job satisfaction whereas only work experience shows some moderating role. Maynard et al. (2006) examine the relationship between various types of underemployment and diverse job attitudes, including job satisfaction, surveying three distinct samples of US employees. They obtain that, although perceptions of underemployment are associated with poor job satisfaction, the relations are not equally strong for all underemployment dimension-attitude facet combinations; that is, relations between underemployment and satisfaction are domain-specific.

These results suggest, as a general conclusion, that over-qualified may feel no more dissatisfied than adequately educated, at least, no in all domains of job satisfaction, and the need of investigating the relationship while considering different facets of job satisfaction. A typical limitation of these studies is that the estimated relationship

between over-qualification and facets of job satisfaction cannot be interpreted causally, since endogeneity may affect (Johnson and Johnson, 2000; Fine and Nevo, 2008; Peiro et al., 2010). We attempt to overcome this by using Instrumental Variable. Before doing that, we describe in Section 3 the data used.

3. Data and descriptive results

The data used in this paper come from the ECVT, a programme, started in 1999 and finished in 2010, produced by the Spanish Ministry of Employment that focuses on employment relationships and, more importantly for our research, on the valuations and attitudes of employees towards their work. The survey addresses employees older than 16, living in households, as being representative of the total employed population, and covering a number of issues relating to working conditions, which allows us to control for a range of individual and job attributes. In particular, we focus on those that have to do with the qualification of individuals and their self-perceived job-match, as well as up to sixteen different domains of job satisfaction. Additional information is provided on socio-demographic variables of employees and on job conditions and attitudes of employees towards work. Overall, it combines objective information on labour, family, and individual characteristics, with pure subjective information on satisfaction with various facets of the job and with the job-match. Micro data are available from the Spanish Labour Ministry since 2001, with the exception of year 2005 when the survey was not carried out. The questionnaire differs throughout the period, with marked differences before and after year 2004. Our focus is then on the most recent period. Specifically, our sample is constructed from pooling the last four consecutive waves, from 2007 to 2010, producing an overall sample of 26,027 individuals.

We have computed over-qualification from a subjective point of view. In this general term we distinguish between over-education and skill under-utilisation. The over-education literature typically considers three ways to measure it: subjective, objective and statistical. An objective measure is based on a comparison between the actual education level and the job-level requirements, established from an evaluation by professional job analysts. The statistical measure of over-education is obtained by comparing years of education attained by an individual with an indicator of the aggregate education level in the occupation in which that individual works. Finally, the subjective measure comes from worker self-assessments. It corresponds to the answer to the question whether the individual feels over- or under-educated for the work they do.

Although subjective measures can be affected by classification errors, it is generally based on all the relevant information (Green and Zhu, 2010) and has more predictive power over the outcome than alternative measures (Johnson et al., 2002; Maynard et al., 2006), such that they are meaningful interpretations of the work environment (Johnson et al., 2002). Feldman et al. (2002) expresses this as subjective over-qualification being a more proximal predictor of employee attitudes and behaviours (see also, Liu and Wang, 2014).

Specifically, we first make use of the worker's responses to the following question.

QUESTION 1 *Do you think that your current job is adequate according to your educational level?*

With the possible answers being

1. *Yes, correct.* We label this as adequately educated
2. *No, below.* We label this as over-educated
3. *No, above.* We label this as under-educated
4. *No, different.* We label this as mismatched.

Since less than 3% of surveyed individuals choose answers 3 and 4, we discard these individuals in our analyses and consider only adequately matched and over-educated.¹ According to Feldman's (1996) classification, this would correspond to the first dimension of underemployment: more education than required by the job.² Table 1 shows the proportion of each of these groups in the total amount.

A second dimension of over-qualification and underemployment is skill under-utilisation. This is more difficult to classify and measure. Here, it is approached by considering a second question in the ECVT:

QUESTION 2 *To what extent is your educational level useful for your job?*

Each individual rates between 0, *not at all*, and 10, *very much*. The answer given to this question may be interpreted as an indicator of skills utilization, since it may well be the case that an individual declares herself as adequately matched and, simultaneously, reports a low degree of usefulness of her studies in her current job. We consider that the portion of the sample rating between 0 and 5 have acquired educational skills that are

¹ Apart from a possible reluctance of individuals to acknowledge being under-educated or mismatched, it is reasonable to consider that experience and on-the-job training may help workers to reduce the self-perception of being under-educated or mismatched.

² The classification includes five dimensions. The second is more experience or skills than required; the third refers to involuntary employment in a field outside of area of education; the fourth, to involuntary employment in part-time, temporary or intermittent work; and the fifth, to underpayment. The two latter are not directly related to qualification and are not studied here.

hardly applicable to their jobs (*non-useful skills*), whereas the half rating 6 or above are thought to make great use of their acquired educational skills (*useful skills*).³

Taking responses to both questions, we can construct a classification of employees according to a self-evaluated mismatch (see Table 1). We define as “properly matched” those who answer 1 to the first question and simultaneously rate 6 or more to the second question. They are neither over-educated nor over-skilled and represent almost two thirds of the whole sample. Those who answer 2 to the first question are labelled as over-educated. We can distinguish between “apparent”, those who report 6 or above to the second questions (over-educated but not over-skilled), and “genuine”, those who report 5 or below (over-educated and over-skilled). The remaining individuals, about 15% of observations, correspond to individuals who report there is not much of a relationship between their realized studies and the tasks they perform at work. These are more difficult to classify; they are differently qualified and we designate them as “unadjusted”. These can approximate a third dimension of underemployment, those who are employed in a field outside of area of education (Feldman, 1996). Whereas responses to *QUESTION 1* allow distinguishing between adequately-educated and over-educated, the combination of responses to both questions permits to partially capture skill heterogeneity among individuals, and distinguish between properly-matched; unadjusted; apparent over-qualified; and genuine over-qualified. Both classifications will be considered in our subsequent analyses.

(Table 1 about here)

Workers are asked a number of questions concerning different aspects of job satisfaction. The overall assessment on job satisfaction at the current job is derived from the answer to the question “*indicate the satisfaction degree in their current (main) job*”, which is rated – on an eleven-point scale - from 0 (*no satisfaction*) to 10 (*very high satisfaction*). Average general job-satisfaction for the pooled 2007-2010 period is rated at 7.29. Only a lower than 20% proportion of respondents admits to a job satisfaction rate below 6. While in many cases (see Johnson et al., 2002; Peiro et al., 2010, among others) the chosen facets of job satisfaction are constructed through confirmatory factor analysis from the answers of individuals in short samples to different items, we use the nationally representative ECVT which directly offers the rates for each of the

³ Other possibilities have been also considered in estimations. Although the percentage of individuals in each group varies, estimated results do not significantly change with respect to those presented in the next sections.

satisfaction dimensions under considerations. The advantages of using large nationally representative samples have been asserted elsewhere (f.i., Liu and Wang, 2012).

Table 2 shows the definition of the different domains considered of satisfaction at the workplace, their corresponding overall average values, as well as distinguishing between over-educated and adequately educated (according to *QUESTION 1*), and across the four groups aforementioned (combining answers to *QUESTIONS 1* and *2*). It can be seen that the general valuation of satisfaction is generally larger than those of the particular domains considered. Whereas many domains show an average value between 6.5 and 7.7, lower values (less than 5) are observed for satisfaction with promotion prospects and with training. Pay satisfaction is near 6 and the lowest value, 3.3, is attached to satisfaction with social benefits or aids provided by the firm. We compute partial correlations between each domain of satisfaction and the general assessment of the job satisfaction finding that they are all positively correlated, but are far from perfect. Thus, the highest coefficients (around 0.60) correspond to satisfaction with activity in work and satisfaction with personal development, and the lowest (between 0.28 and 0.36) satisfaction with hour flexibility and satisfaction with firm-provided social benefits.⁴

On its part, over-educated systematically report lower rates of satisfaction for all domains than those who are adequately educated.⁵ Many differences are about one-scale point, with the two largest exceeding 1.4 (satisfaction with personal development and with training) and the lowest being smaller than 0.5 (satisfaction with working day and vacations). When considering the differences in the set of satisfaction domains between the four types defined of match, in general, properly matched are more satisfied than unadjusted; these more than apparently over-qualified; and these latter more than genuinely over-qualified (values ranging from properly to genuine over-qualified between 1 and 2-scale points). In some cases rates for apparently over-qualified are not lower than those of unadjusted.

4. Methodology and results

4.1 Ordered probit estimation

We follow the standard approach, regressing each domain of satisfaction, and overall job satisfaction, on a range of personal and job characteristics at the individual level,

⁴ Results not shown but available from the authors upon request.

⁵ Median values (not shown) provide an identical conclusion.

adding self-perceived variables capturing over-qualification, according to the following expression:

$$DS_{it}^j = \alpha + \lambda_t + \beta_0 X_{it} + \gamma_0 M_{it} + \varepsilon_{it} \quad (1)$$

where self-reported satisfaction, DS for each domain j , of individual i , in year t depends on the year dummies (λ_t), a vector of individual socio-demographic and job characteristics (X_{it}) and dummies capturing over-qualification (M_{it}).

An important matter referring to our dependent variable is whether the different domains of satisfaction are assumed to be ordinal-interpersonal comparable, or cardinal-interpersonal comparable (Ferrer-i-Carbonell and Frijters, 2004). Interpersonal comparability means that when two respondents give the same answer, they are assumed to enjoy similar satisfaction levels. That is, “individuals have a common understanding of how to translate internal feelings into a number scale, so that numerical values from different individuals are roughly the same” (Ferrer-i-Carbonell and Frijters, 2004: 644). Assuming cardinality supposes that the differences between satisfaction rates are not dependent on the rate itself (i. e. the difference between rating 7 and rating 6 is the same as the difference between scores 3 and 2). In consequence, an individual rating 8, for example, is twice as satisfied as an individual rating 4. In this context, empirical analysis can be done with OLS. However, when ordinality is assumed differences between the rates are not considered to be independent on the rate itself. In this case, an individual rating 8 is more satisfied than another individual rating 4, but the difference is not informative on the relative valuation. The empirical analysis requires hence the use of latent variable models, ordered probit or ordered logit. The assumption of ordinal-interpersonal comparability is then less restrictive and results for ordered probit are now examined in detail.⁶

At the moment we present ordered probit estimates of equation (1) without considering the bias associated with potential endogeneity, to be discussed below. Thus, causality is not investigated and coefficients should be interpreted as only partial correlations. Table 3 shows the results of estimating the relationship between self-perceived over-qualification and the overall concept of job satisfaction while controlling for a set of (observable) personal and job characteristics. Controls that moderate the negative relationship between over-qualification and (domains of) job satisfaction

⁶ We also estimated by OLS. Ferrer-i-Carbonell and Frijters (2004) find that the assumption of cardinality or ordinality does not qualitatively change the results in the studies of happiness, and many non-economic papers rely on OLS estimates since their interpretation is more straightforward. We find little qualitative difference between the results of those approaches. OLS estimates are not reported but are available upon request.

include, in addition to years, gender, educational attainment, occupational category, age (in quadratic terms), nationality, city-population size, working hours, income ranges, type of contract, training at work, activity branches, tenure (in quadratic terms), being in a first job, firm size, unionisation and family variables (marital status, number of child in different age ranges).

The inclusion of earnings is essential. With no market failures, preferences over job amenities would be internalised in the labour market through wages (compensating wage differentials) and one would then not find any separate effect of, say, worked hours or type of contract on any domain of job satisfaction after controlling for income or wages. Nevertheless, studies systematically do find statistical and quantitatively significant effects of various job amenities on job satisfaction as, for example, type of contract (Booth et al., 2012), over-education (Cabral-Viera, 2005), job stability (Origo and Pagani, 2009) or empowerment (Erdogan and Bauer, 2009); thereby contending the competitive view. In other words, any job characteristics that influence utility/job satisfaction may offset, or reinforce, the effect of over-qualification. In consequence, it is necessary to condition the analysis on the characteristics of the job.

Our own results confirm this. The estimates for variables included in our analysis typically follow the standard behaviour in the literature (see Dolan et al., 2008; García-Mainar et al., 2016, for the case of Spain). Age and tenure variables have the typical U shape, indicating that, in the early years, satisfaction declines and then increases. Foreign workers are more satisfied, *ceteris paribus*, than native Spanish workers. Men are less satisfied than women. Higher education and larger population-size cities are associated with lower job satisfaction. The family structure is found to be an important element in shaping job satisfaction: marriage is associated with higher job satisfaction, whereas having infants at home is linked to lower job satisfaction. Regarding work-related variables, working longer hours results in lower satisfaction. Higher income is positively associated with greater job satisfaction. Labour stability and training at work both lead to increases in job satisfaction, with workers being generally more satisfied in non-manual occupations and in services. Being in a first job is positively related to job satisfaction, while firm size and unionisation both reduce job satisfaction.

(Table 3 about here)

Table 4 shows the estimated coefficients of the over-qualification variables for each domain of job satisfaction. The rest of the coefficients are not reported to save space. They show a similar pattern to that described for overall job satisfaction, and results are available from the authors upon request.

INSERT TABLE 4

Focusing on our variables of interest, we consider alternatively the specification distinguishing between over-educated and adequately matched (first column, Table 3 and 4), and the one that simultaneously consider properly matched, unadjusted, apparently over-qualified and genuinely over-qualified (second block of columns, Table 3 and 4). In both specifications results are coincident: over-qualification conveys losses in satisfaction for employees with respect to those who are adequately-educated and properly-matched, respectively. First column in Table 4 shows that over-educated are less satisfied than adequately-educated for every domain of job satisfaction, even after controlling for an ample set of personal, family and job-related variables. When studying different degrees of over-qualification, it can be seen in the next block of columns, that the negative relationship is the strongest in the case of genuinely over-educated followed by apparently over-educated and finally by those unadjusted, concluding thereby that the highest the mismatch, the least satisfied an individual is for any domain of job satisfaction.

On concluding, being in a job requiring less qualification than those held by the employee supposes less satisfaction in all the facets or domains of job satisfaction we have studied. The fact that mismatch is sub-optimal and hence a situation not pursued by employees is immediately derived. Results fit well the relative deprivation theory.

4.2 Addressing endogeneity through IV estimation

When estimating equation (1) we face certain difficulties. The fact that perceived over-qualification and satisfaction, in different domains, are both evaluated subjectively, makes endogeneity biases likely to arise. Whereas results in Table 4 show that there exists a strong negatively relationship between self-perceived over-qualification and measures of satisfaction in different domains, these results may be driven by individual psychological characteristics or personal traits that simultaneously influence in perceived educational mismatch and job satisfaction. Thus, unobserved characteristics that are omitted from the equation, such as self-esteem, anger or boredom, are likely to affect both self-perceived levels of job-match and satisfaction in the workplace. In addition to this problem of common unobserved determinants, reverse causation may be also at work. Individuals who feel more mismatched may become more dissatisfied in the workplace, and then, report lower levels of job satisfaction in different facets; but, also, individual with low levels of satisfaction in certain spheres of the job may report to be more discontented with their qualification/job match. Both possibilities suggest an

endogeneity problem, leading to biased and inconsistent estimates of the causal effect of self-perceived mismatch on the subjective evaluation of facets of job satisfaction.

Since our data are repeated cross-sections, we cannot control for unobserved individual heterogeneity through panel data estimation. One standard approach to account for it is to use IV estimation. This attempts to instrument subjective educational mismatch in order to obtain consistent estimates through 2SLS or GMM. However, the treatment of endogeneity when the dependent variable is ordered is not straightforward. In this context, the two-step method (2SLS) can be viewed only as an approximation of the correct estimator (see e.g. Van de Ven and Van Praag, 1981; Bryson et al., 2004; García-Mainar et al., 2016). A simple way to circumvent this is by assuming that the dependent variable is cardinal. As mentioned above, assuming either ordinality or cardinality of happiness scores has little effect on the qualitative empirical results (Ferrer-i-Carbonell and Frijters, 2004).

We then assume that job satisfaction is cardinal, so that IV estimation can be carried out to control for endogeneity, and tests for exogeneity of the regressors and for the validity of instruments can be routinely used. To aid identification of the effects of interest, we formulate a set of exclusion restrictions. We need to make assumptions about the variables that affect over-qualification but, conditional on these, have no residual impact on each domain of job satisfaction. Specifically, we construct an objective measure of over-qualification through a dummy variable which takes value 1 if the employee has university studies but work in a manual or services unskilled occupations, and 0 otherwise.⁷ Since this measure of over-education is objective, endogeneity biases in the estimation of the relationship between over-qualification and each domain of job satisfaction are (at least) partially controlled for. An equation (2) is added

$$M_{it} = \alpha + \lambda_t + \delta_0 Z_{it} + \delta_1 OQ_{it} + \nu_{it} \quad (2)$$

where Z_{it} is a vector of explanatory exogenous variables included in equation (1) and OQ_{it} is the instrument (either the dummy variable capturing over-education or the ordered variable according to definition 2, whose values are, respectively, 0 properly matched, 1 unadjusted, 2 apparently over-qualified and 3 genuinely over-qualified.) The fitted values for M obtained in this equation are introduced in the estimation of equation

⁷ Manual and services occupations correspond to groups 4, 5, 8 and 9 of the one-digit 1994 *CNO Clasificación Nacional de Ocupaciones* (National Classification of Occupations). This is based on the 1988 ISCO classification, but they do not entirely coincide. Group 4: Clerks; group 5: Services and sales workers; group 8: Machine operators and assemblers; group 9: Elementary occupations.

(1), with estimated values shown in the last columns: estimates for the whole set of coefficients in the case of overall job satisfaction in Table 3 (they show a similar pattern to that in the ordered probit case) and for the coefficients of interest for each domain of job satisfaction in Table 4.

Table 5 reports the results of the tests for exogeneity of the over-qualification variable. When distinguishing between the adequately-educated and the over-educated, the standard Wu-Hausman test rejects the exogeneity of over-qualification in most of the cases (only in overall job-satisfaction, satisfaction with activity, satisfaction with hour flexibility, satisfaction with job stability and satisfaction with training exogeneity is non-rejected), such that inference should be made on coefficients estimated by IV.⁸ The estimated coefficients for over-education, once endogeneity is controlled for (in those cases where it is found that this variable is considered as endogenous), make a marked difference with those shown in first column of Table 4. Now, there are many facets of job satisfaction for which over-educated are not less satisfied than adequately-educated. Only in the dimensions of activity, personal self-fulfilment, autonomy and participation, as well as in the overall dimension, those who are over-educated are significantly less satisfied than adequately matched. In the rest of cases, the relationship is either negative (satisfaction with firm's organisation)⁹ or positive (satisfaction with promotion, satisfaction with relations with management, with workday, with duration of breaks, with paid vacations, with pay, with social support and with health and safety at work), but no longer statistically significant. These results point out at that reverse causation or individual heterogeneity problems may be biasing the estimates, and more dissatisfied workers likely reporting to be inadequately matched.

When the over-qualification instrument/variable is captured by the ordered variable indicating the degree of mismatch, this variable is found to be exogenous –in statistical terms- only in the cases of satisfaction with activity and satisfaction with hour flexibility.¹⁰ This leads to focus on the ordered probit estimates on these two domains only, finding that the relationship between over-qualification and these two domains of job satisfaction is negative. The fact that only in two cases over-qualification can be considered as exogenous, may be suggesting that as the indicator of over-qualification is

⁸ The elaborated objective measure of over-education, that acts as the instrument, is found to be not weak (F-statistics are much above the critical value).

⁹ The other domains in which the relationship is estimated to be negative, satisfaction with hour flexibility, with job stability and with training, the relevant sign is obtained from ordered probit estimation (first column of Table 4) since, in these cases, over-education is found to be exogenous.

¹⁰ The same elaborated objective measure of over-education is used as instrument, and again is found to be not weak in all the cases (see the F-statistics under the IVreg heading).

more precisely measure, the same elements influencing in self-perceived over-qualification, affect the reported value of satisfaction (except in the two cases aforementioned). Dealing with endogeneity using OQ as instrument, we find similar result to those obtained in the case where we distinguish only between adequately-matched and over-educated. The relationship is also negative and statistically significant, in addition to satisfaction with activity and satisfaction with hour flexibility, in the overall domain of job satisfaction and in the facets of personal self-fulfilment, autonomy, participation, and marginally, in satisfaction with training. In the rest of the cases, the relationship is not statistically significant, either positive or negative.

On conclusion, controlling for endogeneity reveals that in many facets over-qualified are similarly satisfied to the adequately-matched suggesting that over-qualification does not necessarily lead to lower levels of satisfaction or sense of deprivation. Whereas in the overall domain of job satisfaction over-qualified are indeed less satisfied, as usually found in the literature, this negative relationship is observed in some few facets such as satisfaction with the activity developed in work, with self-fulfilment, with autonomy at work and with the participation in the decisions related to job tasks. Only in these cases, over-qualified do really feel deprived with respect to workers adequately-matched. In the other domains of job satisfaction under consideration, an over-qualified worker does not feel worse than a properly-matched. Thus, in facets such as relations with management, paid vacations, duration of breaks, workday, job stability or earnings, there are hardly differences in satisfaction between over-qualified and adequately-matched.

If job satisfaction can be considered as a predictor of job performance or voluntary turnover, it seems that the general concept it is in fact negatively related with over-qualification and then individuals may feel deprived. However, in many other facets this is not true and over-qualified workers are as satisfied as perfectly matched workers. Thus, it is not unexpected that, at least some individuals, may get over-qualified not only as a consequence of mismatch due to frictions in job search, job mobility or in other labour market imbalances, but also as a way to obtain some reward compensating disadvantages from over-qualification. Specifically, the evidence obtained in our study indicates that they do not feel less satisfied in earnings, health conditions and safety at work, promotion prospects or job stability, among several others.

Discussion and conclusions

There exists large empirical evidence showing a negative influence of over-qualification on satisfaction which can be rationalised into relative deprivation arguments. However, job satisfaction can be seen as an array of different domains/facets of satisfaction related to performance in the job place. In fact, different studies find that over-qualified workers feel less satisfied in some domains and more satisfied in other domains than adequately matched. For example, if workers over-qualify to obtain a more permanent position or just to escape from unemployment, then they may report a valuation of satisfaction with job stability or job promotion not very different from that reported by an adequately matched worker, counterbalancing lower valuations corresponding to other domains, such as pay or self-fulfilment.

In this paper, we deepen into this argument by hypothesising that, individuals who over-qualify may be not less satisfied than adequately matched if they aim to mask or compensate their lack of skills or to access to employment. Whereas over-qualification is habitually considered as suboptimal, consequence of a mismatch due to search or job frictions, the increasing dispersion in ability and/or skills among equally educated workers may induce individuals to voluntarily acquire more qualifications than those they can productively use in their jobs in order to compensate for a lack in those observed and unobserved skills or as a way of gaining access to employment or to search for a better job. Even more, this signalling role of over-education may be especially important in periods of recession, and/or in areas where unemployment is high, since the decision to invest in education is not only regarded as a way to have access to higher wages but also as a way of insuring against unemployment.

Our results suggest that the relationships between over-qualification and job satisfaction are better measured when considering different domains or facets of the job satisfaction concept. When taking into account simultaneity or reverse causation, the negative relationship seems to remain in the general notion of job satisfaction, as well as in intrinsic facets of job satisfaction such as activity developed, self-fulfilment, autonomy and participation in decisions. The relationship is not negative, overqualified do not feel relatively deprived with respect to adequately matched in extrinsic facets (paid vacations, breaks' duration, workday, job security/stability and earnings) or social relations (relations with management). These results are different to those from Peiro et al. (2010) since they find that three facets (extrinsic, intrinsic and social) are negatively related with over-qualification. Two circumstances may explain these different findings. First, whereas we make use of a national representative sample of Spanish workers evaluation sixteen different domains of job satisfaction, Peiro et al. (2010) use a sample

of individual from three regions of Spain and construct the aggregate facets of job satisfaction from several questions in the survey through confirmatory analysis. Second, and more important, as Peiro et al (2010) acknowledges, the relationship they found cannot be interpreted causally, calling for the need of propensity score or IV techniques. This is what we have done here, so that reverse causation and/or simultaneity bias are dealt with. In fact, our results mimics theirs in that a negative relationship is observed between over-qualification and any domain of job satisfaction when endogeneity is ignored.

In short, whereas over-qualified are less satisfied in the overall dimension of job satisfaction, this is not completely true when considering different domains suggesting that a robust assessment of determinants of job satisfaction implies analysing different facets in the job place. This provides some support to the idea that some degree of over-qualification is voluntarily chosen. It remains to explain which is the mechanism driving this behaviour. The cross-sectional nature of data limits the analysis and forces us to left this for future research.

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Table 1 Definitions and percentages of over-qualification

Question 1 \ Question 2	Non-useful skills (0-5 rate)	Useful skills (6-10 rate)
	26.7%	73.3%
Adequately educated (80.7%)	Unadjusted (18.0%)	Properly matched (62.7%)
Over-educated (19.3%)	Genuinely over-educated (8.7%)	Apparently over-educated (10.6%)

Source: Own elaboration from ECVT 2007-2010.

Table 2 Average job satisfaction across groups

Please, rate your satisfaction with the following aspects in your job place. How satisfied are you with the following job aspects?	Mean	Adequately- educated	Over- educated	Adjusted	Unadjusted	Apparently over-qualified	Genuine over-qualified	Number observations
Job satisfaction	7.29	7.47	6.47	7.62	7.06	6.82	6.05	26,027
Firm's work organisation	6.89	7.05	6.17	7.17	6.80	6.38	5.92	26,027
Promotion prospect	4.99	5.20	4.09	5.49	4.20	4.60	3.47	23,884
Recognition of the work / relations with management	7.09	7.26	6.35	7.38	6.93	6.62	6.02	25,056
Activity in work	7.68	7.87	6.88	8.00	7.49	7.26	6.40	26,027
Personal development / Self-fulfilment	7.42	7.67	6.31	7.81	7.25	6.81	5.70	26,027
Level of freedom/ autonomy at work	7.27	7.43	6.60	7.57	6.99	6.94	6.18	26,027
Participation in decisions related to job tasks	6.57	6.78	5.62	7.02	6.05	6.16	4.95	25,389
Workday /Working day	7.16	7.26	6.68	7.37	6.97	6.93	6.38	26,027
Hour flexibility	6.27	6.41	5.65	6.58	5.89	6.04	5.17	26,027
Duration of work breaks	6.63	6.76	6.07	6.89	6.37	6.37	5.70	26,027
Paid vacations	7.41	7.53	6.87	7.71	6.97	7.23	6.43	26,027
Job stability	7.38	7.54	6.69	7.74	6.93	7.15	6.12	26,027
Firm's provided training	5.73	6.00	4.53	6.37	4.80	5.22	3.67	25,151
Earnings/ pay/ salary	6.01	6.20	5.16	6.39	5.61	5.46	4.79	25,948
Firm's provided social support	3.29	3.43	2.67	3.71	2.47	3.13	2.09	23991
Health and safety at work	7.31	7.43	6.79	7.53	7.13	7.04	6.49	26027

Table 3 Estimates of job satisfaction regression

Job satisfaction	Ordered probit				ivregress			
	Coef.	Std.Err.	Coef.	Std.Err.	Coef.	Std.Err.	Coef.	Std.Err.
Over-education	-0.548***	0.018			-0.639***	0.236		
Unadjusted			-0.224***	0.018			-0.238***	0.086
Apparently			-0.451***	0.022				
Genuine			-0.781***	0.025				
Male	-0.071***	0.015	-0.071***	0.015	-0.107***	0.025	-0.097***	0.025
Education2	0.166	0.182	0.285	0.182	0.265	0.295	0.643**	0.317
Education3	0.179	0.178	0.282	0.179	0.285	0.291	0.632**	0.313
Education4	0.168	0.178	0.264	0.179	0.266	0.293	0.602*	0.314
Education5	0.172	0.179	0.264	0.179	0.250	0.299	0.595*	0.320
Education6	0.186	0.179	0.266	0.180	0.265	0.303	0.567*	0.324
Education7	0.139	0.179	0.236	0.179	0.185	0.304	0.508	0.326
Education8	0.138	0.180	0.220	0.180	0.167	0.312	0.463	0.332
Non-manual	0.058***	0.022	0.021	0.022	0.125***	0.042	0.106**	0.048
Salespersons	0.057**	0.024	0.039	0.024	0.091**	0.039	0.081**	0.040
Labourers	-0.021	0.024	-0.003	0.024	-0.067	0.042	-0.055	0.044
Age	-0.036***	0.005	-0.035***	0.005	-0.058***	0.008	-0.056***	0.008
Age ²	0.046***	0.006	0.045***	0.006	0.073***	0.009	0.071***	0.009
Nationality	-0.104***	0.026	-0.113***	0.026	-0.142***	0.046	-0.129***	0.046
Population2	-0.064***	0.019	-0.063***	0.019	-0.113***	0.030	-0.117***	0.031
Population3	-0.126***	0.023	-0.124***	0.023	-0.210***	0.038	-0.222***	0.038
Population4	-0.121***	0.019	-0.121***	0.019	-0.202***	0.030	-0.216***	0.030
Population5	-0.295***	0.027	-0.296***	0.027	-0.474	0.045	-0.489	0.046
Hours	-0.008***	0.001	-0.008***	0.001	-0.013***	0.001	-0.012	0.001
Income2	0.163***	0.019	0.150***	0.019	0.301***	0.034	0.298	0.036
Income3	0.234***	0.021	0.218***	0.021	0.428***	0.043	0.423	0.045
Income4	0.273***	0.025	0.254***	0.025	0.505***	0.060	0.499	0.061
Income5	0.385***	0.031	0.369***	0.031	0.685***	0.079	0.672	0.078
Permanent	0.178***	0.018	0.170***	0.018	0.303***	0.029	0.300	0.030
Training	0.272***	0.014	0.263***	0.014	0.448***	0.026	0.444	0.027
Industry	0.117***	0.042	0.105**	0.042	0.231***	0.067	0.233	0.068
Construction	0.127***	0.043	0.110**	0.043	0.244***	0.070	0.238	0.071
Services	0.224***	0.041	0.206***	0.041	0.392***	0.066	0.379	0.067
Tenure	-0.013***	0.002	-0.013***	0.002	-0.016***	0.004	-0.016	0.004
Tenure ²	0.020***	0.006	0.019***	0.006	0.019*	0.010	0.020	0.010
First	0.054***	0.016	0.051***	0.016	0.090***	0.028	0.089	0.028
Union	-0.120***	0.017	-0.119***	0.017	-0.221***	0.028	-0.222	0.028
Firm_size2	-0.099***	0.019	-0.098***	0.019	-0.147***	0.031	-0.158	0.032
Firm_size3	-0.131***	0.021	-0.128***	0.021	-0.192***	0.035	-0.188	0.035
Firm_size4	-0.139***	0.019	-0.134***	0.019	-0.208***	0.031	-0.205	0.032
Married	0.064***	0.016	0.063***	0.016	0.105***	0.025	0.097	0.026
Children 0-2	-0.033*	0.019	-0.031	0.019	-0.045	0.032	-0.038	0.032
Children 3-5	-0.009	0.019	-0.007	0.019	-0.009	0.031	0.010	0.032
Children6-14	0.034***	0.012	0.034***	0.012	0.046**	0.019	0.054	0.019
Year 2	0.113***	0.029	0.135***	0.029	0.169***	0.054	0.161	0.055

Year 3	0.204***	0.029	0.215***	0.029	0.269***	0.052	0.250	0.052
Year 4	0.249***	0.029	0.257***	0.029	0.348***	0.052	0.333	0.051
/cut_1_1	-3.165	0.208	-3.184	0.208				
/cut_1_2	-2.971	0.208	-2.987	0.208				
/cut_1_3	-2.723	0.207	-2.737	0.207				
/cut_1_4	-2.434	0.207	-2.444	0.207				
/cut_1_5	-2.138	0.206	-2.145	0.207				
/cut_1_6	-1.569	0.206	-1.571	0.206				
/cut_1_7	-1.098	0.206	-1.096	0.206				
/cut_1_8	-0.448	0.206	-0.443	0.206				
/cut_1_9	0.402	0.206	0.411	0.206				
/cut_1_10	0.946	0.206	0.956	0.206				
LogL	-47599		-47469					
R ² adjusted					0.10		0.11	

Tabla 4. Over-education and Over-qualification coefficients for all facets of job satisfaction

	Ordered probit								IV			
	Over-education		Over-qualification						Over-education	Over-qualification		
			Unadjusted		Apparent		Genuine					
Job satisfaction	-0.548***	0.018	-0.224***	0.018	-0.451***	0.022	-0.781***	0.025	-0.639***	0.236	-0.238***	0.086
Firm's work organisation	-0.377***	0.017	-0.13***	0.018	-0.335***	0.022	-0.493***	0.025	-0.190	0.308	-0.048	0.113
Promotion prospect	-0.335***	0.018	-0.299***	0.019	-0.281***	0.023	-0.544***	0.026	0.437	0.462	-0.149	0.168
Recognition of the work/relations with management	-0.373***	0.018	-0.150***	0.019	-0.321***	0.022	-0.510***	0.025	0.079	0.322	0.059	0.118
Activity in work	-0.512***	0.018	-0.191***	0.018	-0.402***	0.022	-0.748***	0.025	-1.145***	0.245	-0.451***	0.089
Personal development / Self-fulfilment	0.595***	0.018	-0.204***	0.018	-0.473***	0.022	-0.856***	0.025	-2.182***	0.280	-0.823***	0.101
Level of freedom/ autonomy at work	-0.324***	0.017	-0.111***	0.018	-0.270***	0.022	-0.446***	0.025	-1.729***	0.313	-0.632***	0.114
Participation in decisions related to job tasks	-0.394***	0.018	-0.195***	0.018	-0.316***	0.022	-0.590***	0.025	-2.327***	0.374	-0.828***	0.135
Workday/ Working day	-0.221***	0.017	-0.123***	0.018	-0.166***	0.022	-0.348***	0.025	0.461	0.316	-0.184	0.116
Hour flexibility	-0.226***	0.018	-0.170***	0.018	-0.158***	0.022	-0.393***	0.025	-0.353	0.433	-0.125	0.159
Duration of work breaks	-0.233***	0.017	-0.136***	0.018	-0.184***	0.022	-0.359***	0.025	0.544	0.379	0.209	0.139
Paid vacations	-0.246***	0.018	-0.153***	0.018	-0.202***	0.022	-0.373***	0.025	0.098	0.323	0.086	0.119
Job stability	-0.222***	0.018	-0.108***	0.018	-0.169***	0.022	-0.342***	0.025	-0.070	0.319	-0.004	0.117
Firm's provided training	-0.413***	0.018	-0.365***	0.019	-0.331***	0.023	-0.699***	0.026	-0.611	0.409	-0.247*	0.150
Earnings/ pay/ salary	-0.392***	0.017	-0.184***	0.018	-0.359***	0.022	-0.520***	0.025	0.142	0.306	0.086	0.113
Firm's provided social support	-0.22***	0.019	-0.230***	0.021	-0.166***	0.024	-0.399***	0.028	0.503	0.437	0.233	0.162
Health and safety at work	-0.272***	0.017	-0.129***	0.018	-0.213***	0.022	-0.409***	0.025	0.125	0.304	0.059	0.111

Table 5. Exogeneity tests

	LogL oprobit			IV over-education					IV over-qualification					
	Over-edu	Over-qua	R2	Durbin	Wu Hausman	R2	F	Shea's R ²	R2	Durbin	Wu Hausman	R2	F	Shea's R ²
Job satisfaction	-47599	-47469	0.100	1.281 (0.258)	1.279 (0.258)	0.014	376.547 (0.0000)	0.014	0.109	5.031 (0.025)	5.023 (0.025)	0.145	442.002 (0.000)	0.016
Firm's work organisation	-53221	-53183	0.057	4.206 (0.040)	4.199 (0.040)	0.014	376.547 (0.0000)	0.014	0.054	8.951 (0.003)	8.937 (0.003)	0.145	442.002 (0.000)	0.016
Promotion prospect	-52673	-52524	0.073	9.810 (0.002)	9.794 (0.002)	0.014	376.547 (0.0000)	0.014	0.072	18.427 (0.000)	18.404 (0.000)	0.145	442.002 (0.000)	0.016
Recognition of the work/ relations with management	-50480	-50430	0.045	8.510 (0.004)	8.497 (0.004)	0.014	376.547 (0.0000)	0.014	0.040	15.701 (0.000)	15.7680 (0.000)	0.145	442.002 (0.000)	0.016
Activity in work	-47401	-47286	0.081	1.011 (0.315)	1.009 (0.315)	0.014	376.547 (0.0000)	0.014	0.100	0.074 (0.785)	0.074 (0.785)	0.145	442.002 (0.000)	0.016
Personal development / Self- fulfilment	-49970	-49833	0.078	12.554 (0.000)	12.538 (0.000)	0.014	376.547 (0.0000)	0.014	0.112	6.942 (0.008)	6.931 (0.008)	0.145	442.002 (0.000)	0.016
Level of freedom/ autonomy at work	-52136	-52101	0.051	10.800 (0.001)	10.785 (0.001)	0.014	376.547 (0.0000)	0.014	0.069	6.800 (0.009)	6.789 (0.009)	0.145	442.002 (0.000)	0.016
Participation in decisions related to job tasks	-54343	-54251	0.071	12.367 (0.000)	12.350 (0.000)	0.014	376.547 (0.0000)	0.014	0.102	5.227 (0.022)	5.218 (0.022)	0.145	442.002 (0.000)	0.016
Workday /Working day	-52936	-52896	0.030	9.877 (0.002)	9.863 (0.002)	0.014	376.547 (0.0000)	0.014	0.026	15.509 (0.000)	15.490 (0.000)	0.145	442.002 (0.000)	0.016
Hour flexibility	-57426	-57355	0.035	0.615 (0.433)	0.614 (0.433)	0.014	376.547 (0.0000)	0.014	0.037	2.777 (0.096)	2.773 (0.096)	0.145	442.002 (0.000)	0.016
Duration of work breaks	-55739	-55696	0.014	9.967 (0.002)	9.953 (0.002)	0.014	376.547 (0.0000)	0.014	0.011	15.165 (0.000)	15.146 (0.000)	0.145	442.002 (0.000)	0.016
Paid vacations	-51171	-51122	0.105	4.165 (0.041)	4.158 (0.041)	0.014	376.547 (0.0000)	0.014	0.102	9.852 (0.002)	9.837 (0.002)	0.145	442.002 (0.000)	0.016
Job stability	-50105	-50073	0.244	2.157 (0.142)	2.154 (0.142)	0.014	376.547 (0.0000)	0.014	0.245	5.349 (0.021)	5.340 (0.021)	0.145	442.002 (0.000)	0.016
Firm's provided training	-53185	-52938	0.203	1.661 (0.198)	1.658 (0.198)	0.014	376.547 (0.0000)	0.014	0.212	7.266 (0.007)	7.254 (0.007)	0.145	442.002 (0.000)	0.016
Earnings/ pay/ salary	-53891	-53827	0.119	10.524 (0.001)	10.509 (0.001)	0.014	376.547 (0.0000)	0.014	0.114	19.445 (0.000)	19.423 (0.000)	0.145	442.002 (0.000)	0.016
Firm's provided social support	-47724	-46640	0.135	6.298 (0.012)	6.287 (0.012)	0.014	376.547 (0.0000)	0.014	0.128	13.373 (0.000)	13.353 (0.000)	0.145	442.002 (0.000)	0.016
Health and safety at work	-51904	-51860	0.028	5.446 (0.020)	5.437 (0.020)	0.014	376.547 (0.0000)	0.014	0.025	10.089 (0.002)	10.074 (0.002)	0.145	442.002 (0.000)	0.016

Table A1: Variable definitions and average values (2007-2010).

Variable	Definition	Average
Personal characteristics		
Gender	1: man, 0: woman	0.594
Age	Age in years	40.459
Age ² /100	Age squared divided by 100	17.528
Nationality	1: Spanish, 0: foreign	0.872
Education1	Pre-primary	3.15
Education2	Primary	15.56
Education3	Lower Secondary	21.24
Education4	Upper secondary	10.11
Education5	Vocational short	10.50
Education6	Vocational long/	13.28
Education7	Short Bachelor	10.76
Education8	Long Bachelor and above	14.40
Population1	City size lower than 10,000 inhabitants	0.202
Population2	City size between 10,001 and 50,000 inhabitants	0.278
Population3	City size between 50,001 and 100,000 inhabitants	0.120
Population4	City size between 100,001 and 1,000,000 inhabitants	0.317
Population5	City size higher than 1,000,000 inhabitants	0.082
Married	1: live in couple	0.674
Children 0-2	1: children 0-2 years old	0.113
Children 3-5	1: children 3-5 years old	0.108
Children 6-14	1: children 6-14 years old	0.305
Job characteristics		
Hours worked	Number of hours worked per week	39.573
Income1	Up to 1,000 euro per month (net)	0.351
Income2	Between 1,001 and 1,200 euro per month (net)	0.237
Income3	Between 1,201 and 1,600 euro per month (net)	0.228
Income4	Between 1,601 and 2,100 euro per month (net)	0.108
Income5	More than 2,100 euro per month (net)	0.076
Permanent	Permanent contract: 1, fixed-term contract: 0	0.769
Training	The firm has provided some training in the last 12 months	0.446
Agriculture	The employee works in Agriculture	0.035
Industry	The employee works in Industry	0.229
Construction	The employee works in Construction	0.128
Services	The employee works in Services	0.608
Manual	Qualified workers in agriculture, qualified monitoring and monitored workers	0.291
Non-manual	Qualified worker (managers, professionals, technicians, and clerical workers)	0.432
Salespersons	Unqualified workers in services (sales, personal services,...)	0.158
Labourers	Unqualified workers in other branches (labourers, drivers,...)	0.119
Tenure	Number of years in the same firm	8.966
Tenure ² /100	Tenure squared divided by 100	2.318
First	1: This is the first job	0.208
Union	The employee is unionised	0.166
Firm_size1	Firm size lower than 10 employees	0.287
Firm_size2	Firm size between 11 and 50 employees	0.248
Firm_size3	Firm size between 51 and 250 employees	0.168
Firm_size4	Firm size higher than 250 employees	0.297

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